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UNIVERSITY OF  
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## Summer School on Multidimensional Poverty Analysis

3-14 August 2015

Oxford Department of International Development  
Queen Elizabeth House, University of Oxford

*Tabita, Kenya*



*Rabiya, India*



*Stephanie, Madagascar*



*Agatha, Madagascar*



*Dalma, Kenya*



*Ann-Sashia, Kenya*



*Valérie, Madagascar*



## Some Regression Models for the Alkire and Foster Method

**Paola Ballon**

3 – 14 August 2015

Georgetown University, Washington D.C

*Tabita, Kenya*



*Rabiya, India*



*Stephanie, Madagascar*



*Agatha, Madagascar*



*Dalma, Kenya*



*Ann-Saphia, Kenya*



*Valérie, Madagascar*



**Multidimensional Poverty Measurement  
and Analysis: Chapter 10 – Some  
Regression Models for AF Measures**

**Sabina Alkire, James Foster,  
Suman Seth, Maria Emma Santos,  
Jose Manuel Roche, Paola Ballón**

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# Where we are:

**Post-estimation** analyses of  $M_0$  comprise:

- Decompositions into H and A, or by group or region
- Breakdown by dimensions
- Analysis of associations and redundancy.
- Robustness analysis to parameter selection in measurement design.
- Computation of standard errors for statistical inference - confidence intervals and hypothesis testing.
- Analysis of distribution and dynamics over time.

# Outline of the chapter

10.1 Micro and Macro Regressions

10.2 Generalized Linear Models

10.2.1 Classic Linear Regression

10.2.2 The Generalization

10.2.3 Estimation and Goodness of Fit

10.3 Micro Regression Models with AF Measures

10.3.1 A Micro-Regression Example

10.4 Macro Regression Models for  $M_0$  and  $H$

10.4.2 Econometric issues for an empirical model of  $M_0$  or  $H$

# What are we missing?

Indonesia (1993) provides the following **characterisation** (**descriptive**) of multidimensional poverty ( $M0=0.133$ ) (Ballon & Apablaza, 2013)

## MD poor households characteristics of the household head

<i>Average</i>			<i>Proportion</i>	
Years of education	Age	Household size	Male head	Muslim
2.1	25.5	5.1	80%	91%

we **still miss** the « **effect** » (size) of each of these characteristics on overall poverty in a multivariate framework.

# Why is this important?

From a **policy** perspective, in addition to measuring **poverty** we must perform **some vital** analyses regarding the **transmission mechanisms** between policies and poverty measures.

This is to assess how **poverty** is **explained** by non-**M<sub>0</sub>** related variables

# How can we account for this?

Through **regression analysis** we can **account** for the “effect/size” of **micro** and **macro determinants** of multidimensional poverty.

We can differentiate between:

- **‘micro’** regressions: unit of analysis is the household or the person
- **‘macro’** regressions: unit analysis is some “spatial” aggregate, such as a province, a district or a country.

# This lecture

This lecture provides the reader with a **general modelling framework** for analysing the **determinants** of Alkire–Foster poverty measures, at both **micro** and **macro** levels of analyses.

This modelling framework is studied within the class of **Generalised Linear Models** (GLM's).

GLM's are **preferred** as the data analytic technique. They account for the **bounded and discrete** nature of the AF-type dependent variables.

# Micro and Macro Regressions

What are some vital regression analysis we may wish to study with AF measures?

## Micro regressions:

- a) explore the **determinants** of poverty at the household level
- b) create poverty **profiles**;

## Macro regressions

- a) explore the **elasticity** of poverty to economic growth,
- b) understand how **macro variables** such as average income, public expenditure, decentralization, infrastructure density, information technology **relate** to multidimensional poverty **levels** or **changes** across groups or regions—and across time.

# Which are some 'focal' variables to regress?

Dependent variable measure: $Y$	AF	Range of $Y$	Regression Model	Level	Conditional Distribution $p_Y(y)$
Binary ( $c_i \geq k$ )		0,1	Probability	Micro	Bernoulli
$M_0, H$		[0,1]	Proportion	Macro	Binomial

# The classic regression model

$$y_i = \underbrace{E[Y_i | \mathbf{x}_i]}_{\text{deterministic component}} + \underbrace{\varepsilon_i}_{\text{error component}}$$

where:  $E[Y_i | \mathbf{x}_i]$  denotes the conditional expectation of the random variable  $Y_i$  given  $\mathbf{x}_i$ , and  $\varepsilon_i$  is a disturbance or random error.

This model is a **general** representation of **regression** analysis. It attempts to **explain** the **variation** in the **dependent variable** through the **conditional expectation** **without imposing** any **functional** form on it.

# The linear regression model

If we specify a *linear* functional form of the conditional expectation  $E[Y_i | \mathbf{x}_i]$  denoted as  $\mu_{Y_i|\mathbf{x}_i}$

$$E[Y_i | \mathbf{x}_i] = \eta_i = \beta_0 + \sum_j \beta_j x_{ij}$$

we obtain the classic linear regression model (LRM)

$$y_i = \eta_i + \varepsilon_i.$$

$\eta_i$  is referred to as the **predictor** in the generalized linear model.

# Generalised Linear Modelling

The GLM family of models involves **predicting** a *function* ( $g$ ) of the **conditional mean** of a dependent variable as a *linear combination* of a set of **explanatory variables**  $\eta_i$  (**the linear predictor**). This function is referred to as the **link function**.

A GLM takes the form:

$$g(\mu_{Y_i|\mathbf{x}_i}) = \eta_i = \beta_0 + \sum_j \beta_j \mathbf{x}_{ij}$$

Classic linear regression is a **specific case** of a GLM in which the conditional expectation of the dependent variable is modelled by the *identity function*.

# Generalized Linear Regression Models with AF Measures

Dependent variable AF measure: $Y$	Range of $Y$	Regression Model	Level	Conditional Distribution $p_Y(y)$	Link $g(\mu_i) = \eta_i$	Mean function $\mu_i = G(\eta_i)$
Binary ( $c_i \geq k$ )	0,1	Probability	Micro	Bernoulli	Logit $\log_e \frac{\mu_i}{1 - \mu_i}$	$\Lambda(\eta_i)$
$M_0, H$	[0,1]	Proportion	Macro	Binomial	Probit $\Phi^{-1}(\mu_i)$	$\Phi(\eta_i)$

Note:  $\Phi(\cdot)$  and  $\Lambda(\cdot)$  are the cumulative distribution functions of the standard-normal and logistic distributions, respectively. For the binary model, the conditional mean  $\mu_i$  is the conditional probability  $\pi_i$ .

# A binary model in the GLM framework

$$Y_i = \begin{cases} 1 & \text{if and only if } c_i \geq k \\ 0 & \text{otherwise} \end{cases}$$

The outcomes of this binary variable occur with probability  $\pi_i$  which is a **conditional probability** given the explanatory variables:

$$\pi_i \equiv \Pr(Y_i | \mathbf{x}_i) = \mu_{Y_i|\mathbf{x}_i}$$

For a **binary model** the **conditional distribution of the dependent variable**, or random component in a GLM, is given by a **Bernoulli distribution**.

# A binary model in the GLM framework

To ensure that the  $\pi_i$  stays between 0 and 1, a GLM commonly considers two alternative link functions (g): **probit link** - quantile function of the standard normal distribution function, and the **logit link** - quantile of the logistic distribution function.

The logit model (log of the odds) of  $\pi$  gives the **relative chances** of being multidimensionally poor.

$$\log_e \frac{\pi}{1-\pi} = \beta_0 + \beta_1 x_{1i} + \dots + \beta_k x_{ki}$$

# The logit model

$$\log_e \frac{\pi}{1-\pi} = \beta_0 + \beta_1 x_{1i} + \dots + \beta_k x_{ki}$$

The logit model is a linear, **additive** model for the log odds, equation , but it is also a **multiplicative** model for the odds:

$$\frac{\pi}{1-\pi} = e^{\beta_0} (e^{\beta_1})^{x_{1i}} \dots (e^{\beta_k})^{x_{ki}}$$

Our interest lies on conditional mean  $\pi_i$ .

# Interpretation of Model Parameters

The partial regression coefficients  $\beta_j$  are interpreted as **marginal changes** of the logit, or as **multiplicative** effects on the odds.

Thus  $\beta_j$  indicates the change in the logit due to a one-unit increase in  $x_j$ , and  $e^{\beta_j}$  is the **multiplicative effect** on the odds of increasing  $x_j$  by 1, while holding constant the other explanatory variables.

For this reason  $e^{\beta_j}$  is known as the **odds ratio** associated with a one-unit increase in  $x_j$ .

# Example

Poverty profile for West Java, Indonesia in 1993  
(Ballon & Apablaza, 2013)

We regresses the **log of the odds of being multidimensionally poor** (with  $k=33\%$ ) on demographics, and socio-economic characteristics of the household head.

These have been selected on the grounds of **'restraining'** any **'possible' endogeneity issue** that may arise in the construction of this poverty profile.

# Logistic regression results – West Java, 1993

Variable	Parameter Estimate	Robust Std. Err.	t ratio	Significance level	Odds ratio
Years of education of household head	-0.68	0.03	-19.65	***	0.51
Female household head	0.24	0.09	2.71	***	1.28
Household size	0.09	0.01	7.02	***	1.10
Living in urban areas	-0.85	0.07	-11.40	***	0.43
Being Muslim	-0.02	0.32	-0.07	n.s.	0.98

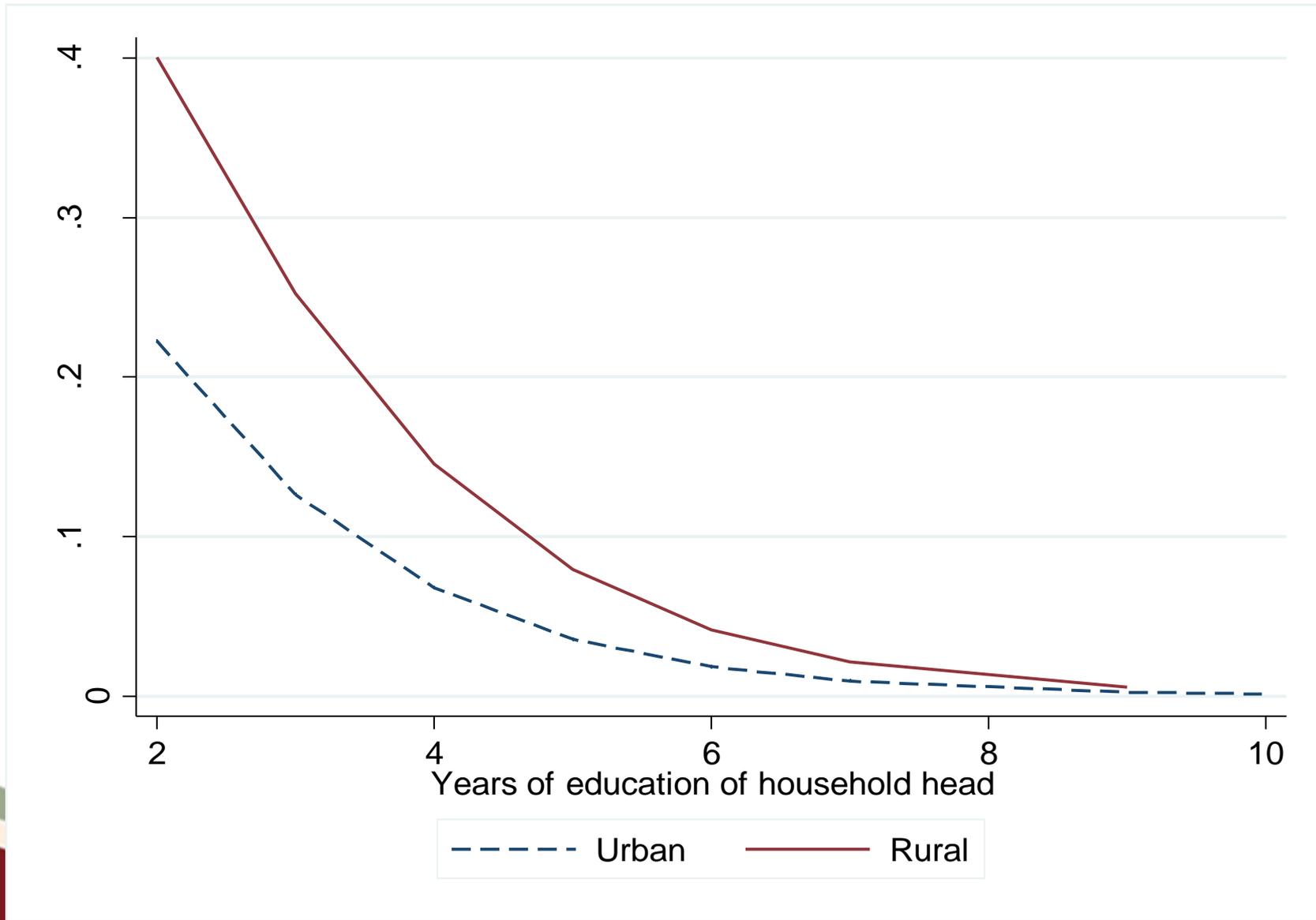
\*\*\* denotes significance at 5% level; n.s. denotes non-significance

Estimated parameters exhibiting a **negative** sign denote a **decrease** in the **odds**, this is obtained as  $(1 - \text{odds ratio}) * 100$ .

For the effect of **education**  $(1 - 0.51) * 100 \downarrow 49\%$ ,

For the effect of **gender**  $(1.28 - 1) * 100\% \uparrow 28\%$ .

# Logistic regression



# Macro Regression Models for $M_0$ and H

H and  $M_0$  are indices, **bounded** between zero and one

Thus an econometric model for these endogenous variables must account for the **shape** of their distribution, which has a **restricted range** of variation that lies in the unit interval.

H and  $M_0$  are therefore **fractional** (proportion) variables bounded between zero and one with the possibility of observing values at the boundaries.

# Papke and Wooldridge (1996) Approach

To model  $H$  or  $M0$  we follow the modeling approach proposed by Papke and Wooldridge (1996).

Papke and Wooldridge propose a particular **quasi-likelihood** method to estimate a proportion.

The method follows Gourieroux, Monfort and Trognon(1984) and McCullagh and Nelder (1989) and is based on the **Bernoulli log-likelihood function**

# The way forward..

Thank you 😊